Cumulative markedness effects and (non-)linearity in phonotactics^{*}

Canaan Breiss^a, Adam Albright^b

^aUniversity of California, Los Angeles ^bMassachusetts Institute of Technology

Abstract

This study uses a series of Artificial Grammar Learning experiments to test for a synchronic relationship between the severity of an individual phonotactic violation and the linearity of its cumulative interaction with other violations, prompted by previous experimental findings (Albright, 2012, Breiss, *resubmitted*). We find that as individual phonotactic patterns are made more exceptionful, their cumulativity moves from linear to super-linear. We evaluate five contemporary phonological frameworks using this data, and find that Maximum Entropy HG, and to a lesser degree Noisy HG, are able to capture the super-linear patterns observed significantly better than other frameworks. Further, we demonstrate that a MaxEnt model provided the same training data as experimental participants exhibits similar emergent super-linear cumulativity, and we explore the weighting conditions under which MaxEnt models yield sub-linear, linear, and super-linear cumulativity.

Keywords: phonotactics, super-linear, super-additive, cumulative constraint interaction, artificial grammar learning, gang effect

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1. Introduction

This paper addresses the relationship between the strength of phonotactic constraints and the way in which multiple coincident violations of such constraints interact in the grammar. Prompted by a dissonance in previous experimental results, we investigate whether there is a causal relationship between the strength of a given phonotactic (as measured in number of exceptions) and how it "gangs" with that of other violations in the grammar. Using an Artificial Grammar Learning (AGL) paradigm similar to Breiss (*resubmitted*), we find that this is the case. As the number of exceptions to a given phonotactic grows, the additional penalty for multiple coincident violations increases beyond what is obtained by simple combination of these independent penalties. That is, participants' acceptability ratings for doubly-marked forms are *lower* than what is obtained by adding up the independent penalties in acceptability for each of those forms' individual violations. We argue that this constitutes a case of *super-linear cumulativity*, and discuss its implications in light of previous experimental and lexical studies of cumulativity.

This finding in hand, we consider what characteristics a phonological framework must possess to capture these effects, and use our experimental data to compare two prominent constraint-based models, Maximum Entropy Harmonic Grammar (MaxEnt; Smolensky (1986); Goldwater and Johnson (2003)) and Noisy Harmonic Grammar (NHG; Boersma and Pater (2008)). We also evaluate a range of other constraint-based theories of phonotactic well-formedness: Stochastic OT (Boersma et al., 1997, 1998; Boersma and Hayes, 2001), Linear Harmonic Grammar (Coetzee and Pater (2008), cf. also the nearly-identical Linear Optimality Theory of Keller (2006)), and the surface-based Maximum Entropy approach (hereafter simply called the Surface-based model (Hayes and Wilson, 2008; Wilson and Gallagher, 2018)). We find that MaxEnt and NHG can capture the relationship between super-linear cumulativity and exceptionfulness substantially better than the other frameworks, and further that MaxEnt outperforms NHG. Finally, we demonstrate that the MaxEnt model can arrive at a weighting which displays super-linear cumulativity when provided only with the training data that participants received.

2. How do speakers compute grammaticality across multiple marked structures?

A growing consensus in the phonological literature supports the view that markedness violations are *cumulative*: when speakers judge the well-formedness

of a word, their judgement is not based on only the most marked structure it contains (as predicted by strict-ranking constraint-based models such as Optimality Theory (Prince and Smolensky, 1993) and its variants). Rather, they attend to all relevant structures and weight their importance according to their severity (as predicted by weighted-constraint models such as Harmonic Grammar (Legendre et al., 1990) and its variants). This aggregation of evidence – termed *cumulativity* (cf. Jäger and Rosenbach (2006) for more on this terminology) - can be observed in two separate domains: that of probability of a given structure in the lexicon, and that of experimentally-determined acceptability or probability.¹ In studies of the contents of lexicons, it has been repeatedly observed that the frequency of word-types which contain a single given marked structure x is greater than that of words which contain both x and also an independently marked structure y. An example of this type of cumulativity can be found in the lexicon of English: as part of a study of English monosyllable phonotactics, Albright (2012) found that 491 (8.2%) of monosyllables in the CELEX database (Baayen et al., 1995) had a stop+l onset, and 47 (3.2%) had a s+stop coda. However, the number of *#stop+l...s+stop#* words was lower than either of these, with only 7 occurrences (0.11%). Experimentally, Pizzo (2015) found that English-speaking participants judged words which violated English syllable-margin phonotactics in one location, ex. *plavb*, *tlag* as less acceptable than one which violated none - plag - plagand crucially more acceptable than those which violated both, ex., tlavb. Breiss (resubmitted) tested for cumulativity in phonotactic markedness using an AGL paradigm, and found that, when trained on a language which conformed to two exceptionless phonotactics, participants judged words that violated both phonotactics as less well-formed than those which violated only one, again demonstrating cumulativity. Kawahara and Breiss (resubmitted) examined cumulativity in sound symbolism, and found that participants combined multiple phonological cues to the same sound-symbolic quality in a cumulative manner.

2.1. Terminology

This paper builds on these experimental and lexical observations of cumulativity to probe the precise relationship between the probability of singly-violating and multiply-violating forms.² Thus, it is important to clearly define terms de-

¹In this paper we use the terms *acceptability* and *probability* interchangeably, though in general this is a non-trivial assumption.

²It is important to note that these types of cumulativity are orthogonal to the distinction between counting and ganging cumulativity (cf. Jäger and Rosenbach (2006)).

scribing the different types of relationship that can obtain.

In general, discussions of constraint cumulativity in probabilistic models of grammar have (often tacitly) assumed that when a word violates multiple constraints, the probability given to the multiply-violating word by the grammar is obtained by multiplying the probabilities of each of the structures it contains. This is not the only relationship possible between the probabilities of singly- and doubly-marked forms, however. For example, consider a language with vowel backness/rounding harmony and consonant nasality harmony, such as the one used in the experiments reported in this paper in section 5. The full range of relations between the probability of a doubly-violating form *poni*, p([poni]), and the probability of minimally-different forms *poti* and *ponu* that violate only one of the two constraints, is laid out in figure 1.³ The diagram groups these types of relations first into those that exhibit any cumulativity (Case 2) from those that do not (Case 1), and then further differentiates the different cumulative ways in which the probability of the doubly-violating form might relate to the singly-violating ones (sub-cases 3, 4, 5).⁴

One possibility - Case 1 - is that the two violations are not cumulative, and thus the probability of the doubly-marked form is equal to least probable of the two individual violations that it contains; such lack of cumulativity is characteristic of the strict domination of Classic Optimality Theory (Prince and Smolensky, 1993). The alternative, Case 2, is a catch-all category that simply states that there is *some* effect of the additional violation on the probability of the doubly-marked form, though what type of cumulativity that is remains unspecified. Cases 3, 4, and 5 are specific types of cumulative interaction, where the probability of the doubly-marked form is, respectively, equal to, less than, or greater than the product of the probabilities of its component violations. Case 3 - linear cumulativity - is what is typically assumed in probabilistic constraint-based grammars.⁵.

³Although this example uses two markedness constraints for simplicity's sake, the concept generalizes to any number of constraints.

⁴Thanks to an anonymous reviewer for inspiring this diagram.

⁵Another set of terminology sometimes used is that of *additivity*, with subtypes *sub-additive*, *additive* and *super-additive*. While there are subtle differences between the two, such as the fact that additivity implies cumulativity but not homogeneity, while linearity implies both, I use the linearity term here in keeping with previous literature (Smith and Pater, 2020, Breiss, *resubmitted*; Kawahara and Breiss (*resubmitted*))



Figure 1: Classes of cumulative relationships.

3. Super-linear cumulativity in experimental data, and how to explain it

The studies demonstrating cumulativity reviewed in section 2 simply served to motivate the claim that phonological grammars allow cumulativity — that is, that we are concerned with determining which sub-cases of Case 2 the phonological grammar allows, rather than being in a Case 1 scenario. Figure 1 makes clear though that the existence of cumulative constraint interactions is not a full answer to what *type* of cumulative constraint interactions the grammar allows.

Using these definitions, we can look at the empirical landscape of cumulativity in a more nuanced way. Recent work by Breiss (*resubmitted*) found using an AGL paradigm that participants inferred linear cumulativity (Case 3) between markedness violations. However, this finding of linearity contrasts with a number of cases of apparent super-linear cumulativity (Case 4) in other lexical and experimental data. For an illustrative example, let us return to Albright's study of monosyllabic English words. In section 1 the example served to illustrate the presence of *any kind* of cumulativity in lexical counts — that words with two marked structures were less common than words with only one. Quantifying what "less common" means, however, reveals that in this instance the cumulativity exhibited is *super-linear* in nature: the independent probabilities of the marked syllable margins alone predicts that $8.2\% \times 3.2\% = 0.22\%$ of the monosyllables in the database — about 16 unique words — should exhibit both the marked onset and marked coda. In fact, however, there are only 7 words which do so, less than half the number predicted by the product of the probability of the marked structures. Thus the relationship between singly- and doubly-marked forms in the English lexicon resembles Case 4, *super-linear cumulativity*, rather than Case 3.

To complement these lexical findings, Albright (2012) replicated a nonword acceptability judgment task from Bailey and Hahn (2001) which asked subjects to rate the acceptability of novel English monosyllables containing onset clusters (e.g. [krɛn, draf]), coda clusters (e.g. [lɛsk, mısp]), or both (e.g. [drısp, krɛsk]). Albright then modeled whether the acceptability of the doubly-marked forms could be predicted solely on the basis of their constituent violations — the expected result in a Case 3 scenario of linear cumulativity — and found that it could not: doubly-marked forms such as [drısp] were rated *less acceptable* than predicted by the sum of their independent penalties, again suggesting Case 4. Other cases of super-linearity have been documented in phonological alternations: for example Smith and Pater (2020) note that super-linear behavior is observed in the interaction of deletion and epenthesis in the surface-realization of French schwa. A full review of super-linear cumulativity in lexical data and alternations is reserved for section 8.3, since the focus of this paper is on static phonotactics.

3.1. In search of an explanatory theory

Although there is growing evidence of super-linear cumulativity in experimental studies, there is not yet a consensus about the synchronic mechanism that produces these effects. A prerequisite for any suitable explanation is that it must be able to explain simultaneously both established examples of linear cumulativity (Case 3, findings of Breiss (*resubmitted*), as well as Durvasula and Liter (*to appear*), and literature reviewed therein) as well as the cases of super-linear cumulativity reviewed here. Below, we outline two theoretical mechanisms which yield testable predictions about how and when we should expect super-linear effects (or lack thereof) in constraint interactions.

3.1.1. The frequency matching hypothesis

Although the studies discussed here are suggestive of a relationship between exceptionfulness and super-linearity in cumulative interactions, the close connection between the lexicon and the phonological grammar prevents us from drawing reliable conclusions from these data. A robust body of evidence, primarily drawn from the study of phonological alternations, indicates that all else equal, speakers will extend statistical trends which hold of their lexicon to novel items in generalization tasks such as wug tests, a propensity which Hayes et al. (2009) call the "Law of Frequency Matching". This tendency poses a confound for the study of super-linear cumulativity using natural languages in light of literature which suggests that the lexicons of many languages exhibit super-linear under-representation (cf. Albright (2012)'s own study on the English lexicon, as well as other cases reviewed in section 8.3). Thus there is an alternative explanation for experimental findings such as those of Albright (2012), namely that they could be an experimental artefact of participants generating acceptability judgements using phonological grammars learned from a lexicon that exhibits super-linear underrepresentation of the very structures which the participants are being asked to judge. In this scenario, the synchronic grammar would not need to encode a relationship between exceptionality and super-linearity, only to allow participants to learn from or draw on the distributional asymmetries in the lexicon, a capacity which has been well established over the past decade or more of phonological research. In this scenario, the experimental findings of Albright and others in section 3 would be due to frequency-matching lexical statistics which exhibit super-linear underrepresentation of doubly-marked forms for reasons independent of the synchronic grammar, while the linear cumulativity for exceptionless phonotactics found by Breiss would reflect the synchronic phonological grammar.

3.1.2. The synchronic hypothesis

In contrast to the frequency-matching account, it could be the case that the synchronic grammar encodes a relationship between the degree of exceptionality of an individual phonotactic, and the type of cumulativity that it exhibits when "ganging up" with other markedness violations. The proposed synchronic mechanism would need to account for both the super-linear cumulativity of exceptionful phonotactics found by Albright, as well as the linear cumulativity of exceptionless ones found by Breiss, and depend on the number of exceptions to the phonotactic in question.

4. Experimental design overview

To distinguish between these two accounts, we used an AGL experiment to test the prediction about the relationship between constraint weight and linearity made by the synchronic hypothesis. If the experimentally-observed super-linear cumulativity discussed above has its source in the synchronic grammar, the linearity of constraint interaction should be sensitive to the number of exceptions to the phonotactics involved (as a proxy for constraint "strength"). On the other hand, if the synchronic grammar plays no role and the results found by Albright (2012); Pizzo (2015) were the result of participants "frequency-matching" their lexicons in acceptability, the number of exceptions to a constraint should not impact the cumulativity that the constraint exhibits.

Breiss (*resubmitted*) used an AGL paradigm to create a "sandbox" environment in which participants were taught a language containing evidence for two phonotactic constraints, free from the confound of the lexicon (cf. section 3.1.1). Participants were then asked to give well-formedness ratings of novel words violating neither, one, or both phonotactics. In Experiment 1 we adopt a very similar experimental design, except that we systematically manipulate the number of exceptions to each phonotactic constraint. The critical manipulation was the training group to which participants were assigned: we created five distinct training Conditions (labeled A-E for convenience), each containing a different number of exceptions to each phonotactic (ranging from 0% (Condition A) to 25% (Condition E) in 6.25% increments). The quantity of interest was the penalty for doubly-marked forms relative to the singly-marked ones, as modulated by the percentage of learning data which did not conform to the majority pattern.

An advantage of pursuing this question using an AGL task rather than studying speakers' intuitions about their native language, as was done in Albright (2012), is that the total number of items in the "lexicon" participants are exposed to is very small. This makes it easier to maintain the subject's ignorance about the language-wide probability of different phonological structures in their environment that would otherwise be compromised by speakers' knowledge of their lexicon, either because they could infer a specific number of doubly-marked structures based on the joint probability of those minority structures in the lexicon, or have noticed an obvious gap in the lexicon corresponding to doubly-marked forms and learned a more complex phonotactic penalizing them directly. Within the tight confines of the artificial lexicon, we can keep the number of doublymarked forms expected via the joint probability of the individual marked structures very close to zero (in the experiment reported below it ranges between 0 and 2 in 0.25-unit increments across the five Conditions). Thus the experiment is in most cases somewhat ambiguous as to whether the nonoccurence of doubly-marked forms is due to linear or super-linear cumulativity, minimizing (though not removing entirely) the possibility that any super-linear cumulativity was overtly learned. Experiment 2 further addresses the question of whether

participants could have been sensitive to the fraction of expected doubly-marked forms absent from the more exceptionful conditions of Experiment 1, despite the sensitivity which would be required to overtly learn super-linear cumulativity from this subtle signal in the data. We find that even when there no super-linear underrepresentation present in the experimental stimuli, learners still display super-linear cumulativity. We take this as strong evidence to support a synchronic connection between the strength of a given phonotactic and the cumulativity it enters into (per the *synchronic hypothesis*), and further speculate that in certain conditions speakers may be biased *in favor* of learning super-linearly cumulative interaction of markedness constraints.

5. Experiment 1: testing for a synchronic relationship between phonotactic strength and linearity

This experiment built upon the results of an experiment reported as Experiment 3b in Breiss (*resubmitted*); data from that experiment is included here as Condition A, the training group whose learning data conformed exceptionlessly to both phonotactics. All data reported below included these participants. Further, the experimental materials were the same as those used in Breiss (*resubmitted*) unless reported otherwise.

5.1. Methods

5.1.1. Participants

375 undergraduate students were recruited from a subject pool at a North American university, and were compensated with course credit. Participants' data were excluded if they failed to meet the criterion for sufficient learning as assessed during the verification phase (n = 0; see section 5.2 for details), for not having spoken English consistently since early childhood (n = 43), and in the case of experimenter error (n = 3), leaving data from 329 participants included in the final analysis.

5.1.2. Stimuli

The exposure phase contained 32 unique CVCV, initially-stressed nonwords, with consonants $\in \{/p, t, m, n/\}$ and vowels $\in \{/i, e, u, o/\}$. One of the two phonotactics was a requirement that consonants harmonize with respect to the feature [nasal], such that both consonants in the word were drawn from either $\{/p, t/\}$ or $\{/m, n/\}$ (exhibiting *nasal harmony*). The other phonotactic required that vowels harmonize with respect to the feature [back], such that both vowels

Condition:		А	В	С	D	E
Percent exceptions to each phonotactic:		0%	6.25%	12.5%	18.75%	25%
Unmarked	potu	32	28	24	20	16
Back exceptions	poti	0	2	4	6	8
Nasal exceptions	ponu	0	2	4	6	8
Doubly-violating	poni	0	0	0	0	0

Table 1: Distribution of stimuli across Conditions in Experiment 1.

in the word were drawn from either $\{/i, e/\}$ or $\{/u, o/\}$ (*backness harmony*). For more on these types of consonant and vowel harmony respectively, see Hansson (2010); Walker (2011).

Five distinct training Conditions (A-E) were distinguished by the number of items that violated each of the phonotactic patterns in the language: 0%, 6.25%, 12.5%, 18.75% or 25%. There were no training items which violated both phonotactics at once, so even in the most exceptionful Condition (Condition E) each phonotactic received support from 75% of the words in the training phase. Table 1 below displays the counts and violation profiles of stimuli.

The verification phase used 16 pairs of minimally-differing nonwords: one member of each pair was a fully-conforming word from the exposure phase, and the other was created by reversing the featural specification for backness (and rounding) or nasality of one of the consonants or vowels in the fully-conforming word. This yielded a pair of words differing only in a single instance of that phoneme. 8 pairs differed in a violation of nasal harmony, and 8 in violation of backness harmony, with differences between pair-members balanced for segmental placement and identity. Verification pairs were balanced so that when a fully-conforming verification word had identical consonants (ex. *totu*), it differed only in the violation of backness harmony (ex., *totu* vs. *toti*). The same condition was imposed on verification trials whose conforming word contained identical vowels. There were no doubly-violating words in the verification phase, since its purpose was simply to ensure that participants had learned each of the two phonotactic constraints independently.

The test phase used a set of 48 novel nonwords which varied in conformity both phonotactics. 24 conformed to both phonotactics (ex. *potu*), eight violated only the nasal-harmony phonotactic (ex., *ponu*), eight violated only the backness-harmony phonotactic (*poti*), and eight violated both the nasal-harmony and backness-harmony phonotactics (*poni*).

All words were recorded in a sound-attenuated room by a phonetically trained

female native English speaker using PCQuirer. They were digitized at 44,100 Hz and normalized for amplitude to 70 dB.

5.2. Design

Participants were assigned to one of the five Conditions, and learned the language by listening to a continuous speech stream containing 20 randomized repetitions of the 32 words selected for that particular training phase. After exposure, participants completed 16 self-paced two-alternative forced choice verification trials. Participants were allowed to advance to the generalization phase if they learned each of the phonotactics to a non-significantly-different degree. This was operationalized by imposing a condition that the difference in number of correct answers between pairs differing only in a nasal harmony violation and those differing only in a backness harmony violation was not allowed to be greater than 2, chosen by using Fisher's exact test (Fisher, 1934) to determine the level at which the proportion of correct answers for each phonotactic significantly differed, across the range of possible accuracies. If participants did not meet criteria after two exposure blocks (one initial and one after failing to meet criterion during the verification phase), they were simply asked to complete the final demographic questionnaire and did not generate data in the generalization phase (although recall from in section 5.1.1 that no participants were excluded for this reason).

If participants met criteria on the verification phase, they advanced to a generalization phase which consisted of a ratings task containing 48 novel words in which participants were asked to rate each of the words on a scale from 0 (*very bad*) to 100 (*very good*) based on how good they sounded as an example of the language they had learned during the exposure phase. At the end of the experiment, demographic and language-background information was collected. The entire experiment lasted approximately 20-30 minutes, depending on the number of additional exposure blocks each participant required.

5.2.1. Procedure

The experiment was conducted in a sound-attenuated room using a modified version of the Experigen platform (Becker and Levine, 2010). At the start of the experiment, participants were informed that they would first be learning a new language, and that they then would be tested on their knowledge of that language. During the exposure phase, participants were instructed to simply sit and listen to the speech stream and, if they felt themselves getting bored, to try to count how many unique words they could find in the speech stream (this task was suggested simply to encourage participants to attend to the speech stream). The exposure phase lasted about ten minutes.

Following the exposure phase, participants completed a self-paced verification phase. On each verification trial participants were played a pair of nonwords in a random order, and were instructed to choose the one that sounded like it could belong to the language they had learned. The generalization phase followed a similar structure, except that each trial containing a single novel nonword to which participants assigned a numerical rating. After completing the generalization phase (or after failure to meet criterion during the verification phase), participants completed a brief demographic questionnaire.

5.3. Analysis

To determine the effect of the number of nonconforming items in training on the linearity of the constraint interaction, we fit a linear mixed effects regression model using the 1me4 package (Bates et al., 2015) in R (Team et al., 2013) to the ratings data from the generalization phase. The model included as fixed effects a three-way interaction between two binary predictors (violation of vowel harmony (y/n, reference level = n), violation of consonant harmony (y/n, reference level = n)) and a continuous fixed effect corresponding to the percentage of exceptions to individual phonotactics in a given participants' training condition (hereafter simply "Condition"), as well as all subsidiary two-way interactions and main effects. Analysis began by fitting a maximally-specified model (following Barr et al. (2013)) and simplifying as necessary to achieve convergence. The final model contained the fixed effects outlined above, plus random intercepts for participant and nonword.

5.4. Results

The results of the experiment are plotted in figure 2. Statistical analysis revealed that violating the nasal-harmony phonotactic was associated with significantly lower ratings ($\beta = -24.93$, p < 0.001). The interaction between violation of nasal-harmony and Condition was significant ($\beta = 0.29$, p < 0.001), indicating that as the percentage of forms violating the nasal-harmony phonotactic in the training data increased, novel forms which violated this phonotactic were judged less ill-formed. The analogous main effects and interaction between violation of the backness-harmony phonotactic and training group was also significant (main effect: $\beta = -9.95$, p = 0.015; interaction: $\beta = 0.19$, p < 0.001). There was also a significant main effect of training group, indicating that as as the number



Figure 2: Experiment 1 results, group-level rating plotted on the vertical axis with standard error, Condition plotted on the horizontal axis. Color denotes which phonotactics were violated.

of fully-conforming words heard in training decreased, fully-conforming words were judged less well-formed as a baseline ($\beta = -0.18$, p < 0.001).

Critically, the three-way interaction between violation of nasal-harmony, violation of backness-harmony, and Condition was significant ($\beta = -0.17$, p < 0.002). As the percentage of nonconforming words in training increased, the difference between singly-marked and unmarked items *decreased*, while the relative markedness associated with the doubly-marked items remained approximately unchanged. Interpreting these findings in light of the different cases of cumulativity in figure 1, we can see that in Condition A, which probed the cumulativity of exceptionless phonotactics, the rating given to the doubly-violating form (purple) is below the nasal-violating form (teal) by approximately the same amount that the backness-violating form (green) is below the non-violating form (red). This is the intuitive correspondent of the two main effects of violating each phonotactic, and the non-significant two-way interaction between them — linear cumulativity. Moving rightward along the horizontal axis to Condition E we see that while there is effectively no penalty for violating backness harmony. Put

another way, the cumulativity of backness and nasal harmony (purple) is superlinear because the rating assigned to it is less than the sum of the differences between non-violating (red) and backness-violating (green) — essentially zero and between non-violating and nasal-violating (teal).

5.5. Local discussion

Experiment 1 tested for a causal link between the strength of a given phonotactic in isolation and the linearity of cumulative constraint interactions in which it participates. We found that learners judged doubly-violating items as more ill-formed than expected when trained on a language which contained more exceptions to individual phonotactics than when trained on a more categorical language, supporting such a connection.

5.6. Against an under-learning explanation

Because the fully-conforming forms became indistinguishable in generalization from those violating only the backness-harmony phonotactic at the highest level of exceptionality (Condition E), it is possible that participants simply *did not learn* the backness harmony phonotactic in this Condition. We deem this unlikely, because participants were required to learn both phonotactics to nonsignificantly-different degrees, per the verification phase. Despite this, it may be that a preponderance of participants were simply skating through at chance on the backness phonotactic and above chance on the other.

To test for this possibility, we calculated each participant's *nasal advantage* score, a measure ranging between -2 and 2 which corresponded to the difference between the number of correct answers (out of 8) that participant gave on questions testing backness- vs. nasal-harmony in the verification phase. A positive score indicates that a participant got more correct answers on the nasal-harmony-assessing questions, and a negative score indicates the reverse. If participants were simply not learning the backness-harmony phonotactic, we should expect to see participants in training Conditions wither more exceptions having a higher *nasal advantage* score. Figure 3 plots nasal advantage scores by Condition, with a linear trend.

A linear model confirmed the visual impression that training Condition (coded as a numerical predictor corresponding to the percentage of training data conforming to both phonotactics) does not significantly predict nasal advantage score ($\beta = -0.015$, p = 0.791). We therefore deem it unlikely that the effects observed in Experiment 1 are due to *insufficient learning* of the backness harmony phonotactic.



Figure 3: Nasal advantage score by Condition: one dot is one participant's score (jitter added for readability).

5.7. Was super-linear cumulativity emergent or overtly learned?

Recall that as the number of singly-violating forms in training increased across Conditions, the number of doubly-violating forms in training remained zero. This design maintained a "sandbox" environment, ensuring that generalization to doubly-violating forms was unaffected by the broader lexical statistics of participants' native language(s). A side-effect of this choice, however, was that the training data contained subtle super-linear under-representation of doubly-marked forms, raising the possibility that the super-linear cumulativity observed in participants' responses was explicitly learned from the data, as anticipated in section 4. Put another way, the more individual exceptions to each phonotactic a participant sees, the more conspicuous the absence of a doubly-violating form becomes. In the most exceptionful training phase, Condition E, participants would expect to hear 2 doubly-violating tokens of the form *poni* based on the product of the probability backness-harmony violations and nasal-harmony violations in their training data (using a formula analogous to of Case 3 in figure 1). Instead they heard zero — a blatant case of super-linear underrepresentation of doubly-marked forms, exactly of the type discussed in section 3.1.1, albeit a much more subtle one than found in natural language lexicons (Albright, 2008; Yang et al., 2018; Albright, 2012).

Thus the results of Experiment 1 are ambiguous between supporting a synchronic grammar which incorporates a relationship between exceptionality and linearity of cumulativity that allows super-linear cumulativity to "emerge" under conditions with more exceptionful stimuli, or simply indicating that learners are sensitive enough to infer a conjoined constraint on the basis of (maximally) 2 missing forms in their learning data (cf. Shih (2017) for more on the role of constraint conjunction in weighted-constraint grammars). We conducted a followup experiment to distinguish between these two possibilities.

6. Experiment 2: a control for Experiment 1's Condition E

To determine whether the super-linear cumulativity observed in Experiment 1 was emergent or overtly learned, we carried out a replication of Condition E from Experiment 1, except that the training data included two doubly-violating forms, thus removing their super-linear underrepresentation in the training data. If participants in Experiment 2 do not exhibit constraint cumulativity which significantly differs from that exhibited in Experiment 1 Condition E, we can conclude that the dependency in the synchronic grammar between exceptionality and super-linear cumulativity must exist in the synchronic grammar, and at least in certain cases must be strong enough to actually override speakers' tendency to frequency-match their input, leading to non-veridical representation of superlinear cumulativity in the face of non-super-linearly-underrepresentative learning data (cf. Kiparsky (1982)). On the other hand, if participants exhibit linear cumulativity in this experiment, we can conclude that the participant is sensitive enough to detect such patterns in the learning data, suggesting support for the possibility that speakers are highly attuned to the nuances of their lexicon which exhibits super-linear underrepresentation (cf. section 3.1.1).⁶

⁶Note that even if participants do not distinguish between Experiment 1's Condition E and Experiment 2, it may still be possible for speakers to overtly learn a super-linear pattern in their data; all we will have demonstrated is that *in this case* the super-linear cumulativity in Experiment 1 was not due to overt learning.

Experiment:		1E	2
Unmarked	potu	16	16
Back exceptions	poti	8	8
Nasal exceptions	ponu	8	8
Doubly-violating	poni	0	2

Table 2: Distribution of training items by type, comparing Experiment 1 (Condition E) to Experiment 2.

6.1. Methods

6.1.1. Participants

86 undergraduate students were recruited from the same subject pool to participate in the experiment, of which 15 were excluded for not being native speakers of English, leaving data from 71 participants for analysis. The current experiment only had one training Condition which was identical to that used in Experiment 1's Condition E, except that two of the singly-violating forms were altered so as to also violate the other phonotactic; see table 2. Compensation, design, and procedure were identical to those of Experiment 1.

6.2. Analysis

To test whether the linearity of cumulativity exhibited in Experiment 2 differed meaningfully from that of Condition E of Experiment 1, the two datasets were analyzed together in a mixed-effects linear regression model. In contrast to Experiment 1, we used a Bayesian implementation of the model so as to allow for direct interpretability of a possible null result, using the brms package (Bürkner et al., 2017).⁷ Bayesian models estimate a range of probable values for the parameters of interest; thus we can conclude that an effect is robust to the extent that 95% of these values, a measure known as a 95% Credible Interval (abbreviated to "95% CI", followed by upper and lower bounds in square brackets), does not include zero. The inverse of this is that if the range is centered on zero, then we can say there is no evidence of an effect for the parameter of interest. For a linguistically-oriented introduction to Bayesian methods for both

⁷The model we fit used default weakly-informative priors, with a burn-in period of 1000 iterations followed by a sampling period of 1000 iterations. We ran four chains to ensure thorough exploration of the posterior distribution, and all \hat{R} values were between 1 and 1.01, indicating that the chains mixed successfully.

theory-building and data analysis, see Nicenboim and Vasishth (2016); for tutorial materials on the brms package in a linguistic context, see Nalborczyk et al. (2019); Vasishth et al. (2018); for a more general primer in Bayesian statistical modeling, see Kruschke (2014).

As in Experiment 1, the dependent variable was the numerical rating given to each word in the generalization phase. Also as in Experiment 1 the model contained a fixed effect of whether the form violated backness harmony (y/n,reference level = n), whether the form violated nasal harmony (y/n, reference level = n), and a binary factor for Experiment (*one/two*, *reference level* = *one*), as well as all two- and three-way interactions of these predictors. The model also contained random intercepts for nonword with slopes for Experiment, and random intercepts for subject with slopes for the interaction of the two binary phonotactic predictors.

We can interpret the results of the model as follows: if the 95% Credible Interval for the three-way interaction of violating backness harmony, violating nasal harmony, and Experiment excludes zero, it indicates that the degree of linearity in the cumulative interaction of violating both phonotactics together compared to their independent violations differed meaningfully between studies. If the 95% Credible Interval for the interaction is centered on zero, we can conclude that the cumulative effect of violating both phonotactics did not differ between studies, and thus was unlikely to have been overtly learned in Experiment 1.

6.3. Results

The results of Experiment 2 are shown in figure 4. Violating nasal harmony resulted in lower ratings ($\beta = -12.72$, 95% CI [-21.49, -3.85]), while the effect of violating backness harmony did not ($\beta = -0.10$, 95% CI [-8.91, 8.72]). One experiment was not reliably associated with higher ratings than the other overall ($\beta = 3.89$, 95% CI [-0.87, 8.35]). The coefficient for the interaction between violating backness and nasal harmony did not differ meaningfully from zero ($\beta = -8.87$, 95% CI [-22.44, 4.96]), nor did the coefficient for the interaction between Experiment and violating backness harmony ($\beta = -0.72$, 95% CI [-4.50, 2.96]), nor did the coefficient for the interaction between Experiment and violating nasal harmony ($\beta = 2.56$, 95% CI [-2.30, 7.22]). Turning to the quantity of interest, the credible intervals for coefficient of the three-way interaction between violating backness harmony, violating nasal harmony, and Experiment surrounded around zero ($\beta = 1.89$, 95% CI [-4.13, 7.73]).



Figure 4: Comparison of mean and standard error of ratings by word type in Experiment 1 Condition E, and Experiment 2.

6.4. Local discussion

Experiment 2 tested for whether the super-linear cumulativity observed in Experiment 1 was a result of participants overtly learning a super-linear penalty from the super-linear underrepresentation in their data. We found that the linearity of cumulativity was not affected by whether or not the training data contained a subtle super-linear pattern, a possible confound to the results of Experiment 1. We take this to be compelling evidence in support a synchronic link between exceptionality in learning data and super-linear cumulativity, as discussed in section 3.1.1, and against the possibility of the effect having been overtly learned. Further, the fact that participant's responses contained superlinear cumulativity in spite of the fact that doing so directly contradicted their learning data suggests that under certain circumstances, learners may be biased *towards* super-linear patterns. Much further experimental inquiry on this topic is needed to expand and verify this claim, and it is thus left for future work. With this data in hand, we turn to how these data can be accounted for by phonological theories.

7. Locating super-linearity in phonological models

Since we observed that increasingly super-linear cumulativity emerges from increasingly exceptionful learning data in a context free from the confound of the lexicon, we can ask what qualities a phonological theory must possess in order to capture this dependency. Two theories which explicitly encode non-linear relationships between Harmony and probability are MaxEnt and NHG models; for in-depth discussion, see (Smith and Pater, 2020; Zuraw and Hayes, 2017; Hayes, 2017, 2020; McPherson and Hayes, 2016, Kawahara and Breiss, *re-submitted*; Hayes, *in progress*). The schematic diagram in figure 5 depicts the sigmoidal relationship of MaxEnt, which we use here for illustration. NHG and Linear HG also have such curves which differ in subtle ways, which will not be the focus of discussion here. Rather, the focus below is on demonstrating that in such theories, the same change in Harmony can translate to different changes in probability.



Figure 5: A schematic illustration of a sigmoidal relationship between Harmony (horizontal axis) and probability (vertical axis). The red, green, and blue lines are tangent to the curve, and demonstrate super-linear (red), linear (green), and sub-linear (blue) cumulativity.

The horizontal axis counts violations of a generic scalar markedness constraint, and the vertical axis plots the probability of a candidate violating it competing against a candidate with a single violation of a binary faithfulness constraint (this schematization is heavily inspired by Zuraw and Hayes (2017); Hayes (2020)). As the number of scalar constraint violations increases, the probability of the candidate violating the scalar markedness constraint decreases in a nonlinear fashion. In this scenario, the relationship between different numbers of violations of the scalar markedness constraint are non-linear, technically reminiscent of counting cumulativity (cf. Jäger and Rosenbach (2006) on this terminology). To see how the same mechanism applies in cases of ganging cumulativity, we need only consider the case where the horizontal axis plots the Harmony of a candidate which is a composite of contributions from several different markedness violations. Here the non-linearity of the relationship between Harmony and probability is evident in that the same amount of Harmony penalty yields differently-sized decrements in probability at different places on the horizontal axis. This corresponds to the different slopes of the tangent lines in figure 5. Below we compare a variety of phonological frameworks on their ability to model the experimental data.

7.1. Two questions to ask when modeling experimental data

Phonological theories can be evaluated on their ability to model experimental results in (at least) two ways: one is to ask whether, given the experimental results, a model can achieve a good fit to the experimental data. Success on this metric indicates that the empirical data are in the space of grammars which are allowed by the framework in question — that is, the framework in question is *descriptively adequate* with respect to the phenomenon, and the model needs no further aid from linking hypotheses to capture the data. Given that a framework *can* capture the experimental results, a more rigorous criterion is *explanatory adequacy*: the proposed model should reproduce the crucial data pattern in the experimental findings given the same training data as the experimental subjects. In a sense, this is asking the model to "take the experiment" itself. Below, we evaluate MaxEnt and NHG using these criteria.

7.1.1. Model setup and fitting

We adopt a linking hypothesis which interprets the experimental rating (0-100) as the probability of endorsement in a binary lexical decision task (0%-100%). Therefore, we employ a model structure in which each attested form competes with a single alternative candidate, the Null Parse (Prince and Smolensky, 1993; Albright, 2008, 2012). An example of one such competition is in table 3.

/poti/	AGR([back])	MParse	H
	2.5	1	
[poti]	1		2.5
i™ null		1	1

Table 3: Competition with the null parse in a schematic tableau.

Our choice to model the experimental data as competition with the Null Parse is motivated the fact that it is in competition with another candidate that MaxEnt and NHG exhibit a sigmoidal relationship between probability and acceptability, and thus we give the models the best chance to succeed. Although often used in cases of phonologically-conditioned gaps in morphological or syntactic contexts (Rice, 2005; Prince and Smolensky, 1993; McCarthy, 2005, among others), in the context of phonotactics the Null Parse model can be thought of as setting a minimum threshold of well-formedness as a criterion of existence. In the context of a probabilistic model of phonotactics, this threshold is "soft", allowing exceptions but dramatically reducing likelihood of forms with a Harmony penalty greater than the weight of MPARSE. Despite a superficial similarity to traditional faithfulness-based approaches to phonotactics (Prince and Smolensky, 1993; Tesar and Smolensky, 1998; Prince and Tesar, 2004), in the Null Parse model the phonotactic grammar allots a probability to a possible form, the alternative to which is nonexistence (rather than repair), as in the markedness-only approach to phonotactics taken by Hayes and Wilson (2008); Wilson and Gallagher (2018) and others.

7.2. Evaluating descriptive adequacy

We first evaluated the descriptive adequacy of MaxEnt and NHG to capture the results of Experiment 1. Figure 7.2 provides predicted ratings derived from the statistical model in section 5.4.⁸

For MaxEnt and NHG, we fit three parameters for each Condition: the weights of MPARSE, AGREE([nas]), and AGREE([back]).⁹ Figure 7.2 displays the experi-

⁸We used model-predicted group means per Condition instead of simply averaging over ratings for each category so as to normal potential by-participant and by-item idiosyncrasies in the experimental results.

⁹The MaxEnt models were fit using the Solver utility in Microsoft Excel (Fylstra et al., 1998). The NHG models were fit using OTSoft (Hayes et al., 2003), with the following settings: 1,000,000 learning iterations, 100,000 testing iterations, initial plasticity = 0.01, final plasticity = 0.001;

		Cond	lition			
Source	Phonotatics violated	A	В	С	D	E
Experiment 1	None	0.70	0.67	0.68	0.64	0.62
	Nasal	0.44	0.52	0.50	0.54	0.49
	Backness	0.61	0.61	0.64	0.62	0.62
	Backness and nasal	0.35	0.40	0.38	0.44	0.40
MaxEnt	None	0.70	0.70	0.70	0.66	0.64
	Nasal	0.44	0.51	0.48	0.52	0.47
	Backness	0.61	0.60	0.62	0.60	0.60
	Backness and nasal	0.35	0.41	0.40	0.46	0.42
NHG	None	0.77	0.78	0.79	0.80	0.76
	Nasal	0.43	0.53	0.55	0.55	0.51
	Backness	0.51	0.53	0.55	0.55	0.51
	Backness and nasal	0.29	0.36	0.38	0.39	0.35

mental findings alongside the best-fitting predictions of the MaxEnt and NHG models; these values are also given in table 7.2.

Table 4: Mean group-level percent endorse for Experiment 1 (converted from predicted ratings by dividing by 100, obtained from the regression model in section 5), and best-fitting predicted values from MaxEnt and NHG grammars.

The two sets of experimental findings and model predictions were then compared by calculating the sum, maximum, and mean of the squared error between the predicted and observed data; these are given on table 5.

Model	Sum squared error	Mean squared error	Max squared error
MaxEnt	0.0057	0.0003	0.0008
NHG	0.1264	0.0063	0.0240

Table 5: Sum, mean, and maximum squared error for the MaxEnt and NHG models.

noise applied by-constraint.



Figure 6: Attested (left) and model predicted (MaxEnt, center; NHG, right) experimental results.

It is clear from both figure and table 5 that while the MaxEnt model achieves a qualitatively and quantitatively tight fit to the experimental data, the NHG matches its training data with less fidelity. Why should this be? As noted in Smith and Pater (2020), although NHG is able to capture super-linear cumulativity, its ability to do so is more limited than MaxEnt. We demonstrate in the next section, however, that both the MaxEnt and NHG models fare significantly better than other models.

7.3. Evaluating other models

To this point we have considered only those frameworks which are known to have a sigmoidal relationship between Harmony and probability — MaxEnt and NHG — as candidate models of the phonological grammar. There remain, however, a number of other phonological frameworks commonly in use which have not been considered, Stochastic OT (Boersma et al., 1997, 1998; Boersma and Hayes, 2001), Linear Harmonic Grammar (Coetzee and Pater (2008), cf. also the nearly-identical Linear Optimality Theory of Keller (2006)), and the surfacebased Maximum Entropy approach (Hayes and Wilson, 2008; Wilson and Gallagher, 2018). So as to avoid an abundance of theories being underdetermined by a dearth of discriminating data, we briefly take up these models here. We demonstrate that these models fare substantially worse on the measure of descriptive adequacy, using the same linking hypothesis as in section 7.1.1.

7.3.1. Model setup and fitting

The Stochastic OT model was fit using OTSoft with the same settings as the NHG model. Fitting the Linear HG model followed the protocol for the MaxEnt model described above, except that the negative Harmony of each candidate was not exponentiated in the calculation of the probability of occurrence. This reduces each binary competition to an instance of the Luce Choice Axiom (Luce, 1959), as previously noted by Pizzo (2015). The surface-based model necessitated a slightly more involved approach: because the it yields a probability distribution over candidates which sum to 1, each of the five separate MaxEnt models without the Null Parse candidate contained four candidates, corresponding to a fully-conforming stimulus (potu), two singly-violating stimuli (poti, ponu) and a doubly-violating stimulus (poni). Constraints and violations were assessed as in the models discussed in section 7.1.1, with the exception that the markednessonly models contained no MPARSE constraint because the Null Parse was not a candidate. Finally, to allow direct comparison between the model output and Experiment 1, the probability assigned to each word type was raised to the power of 1/T, where T is a free *temperature* parameter.¹⁰ The models were fit in Excel using Solver, with the objective function of maximizing the sum of the log-likelihood of the data (by allowing the constraint weights to vary, as is standard in Max-Ent) plus the negative sum of the squared error between the model's predicted probabilities for each category and those obtained in Experiment 1 (with T free to vary to allow for a best-fit scaling parameter for each Condition's model).

These procedures yielded the following fits, plotted in figure 7 and displayed in table 6.

Summary statistics for these models, as well as the best-performing MaxEnt model from table 5, are shown in table 7.

As anticipated, all of these models fall substantially short of even the lesswell-fitting NHG model on the measure of descriptive adequacy.

¹⁰For prior usage of *temperature* as a scaling factor in comparing model-generated and experimentally-obtained data, see Hayes and Wilson (2008), Breiss (*resubmitted*).



Figure 7: Attested (left) and model predicted (Stochastic OT, center left; Linear HG, center right; Surface-based model, right) experimental results.

7.4. Evaluating explanatory adequacy

Since we have observed that the experimental data are approximated well by the Null Parse-equipped MaxEnt model, we can ask whether the relationship between exceptionality in learning data and super-linear cumulativity emerges from the framework given the same training data as the human participants, as suggested in section 3.1.2.¹¹ We fit a set of 5 MaxEnt grammars, providing each with the training data of one of the Conditions in Experiment 1 (as in table 1). Each grammar was allowed to set the weights of MPARSE, AGREE([nas]), and AGREE([back]) to match the frequencies in training data as closely as possible this corresponds to maximizing the likelihood of the training data, as is standard with MaxEnt models. However, because participants exhibited markedly different responses to violations of AGREE([nas]) and AGREE([back]), we additionally required the model to match the average difference in ratings between nasalharmony and backness-harmony violators across Conditions, as measured using the sum standard error, by modifying the likelihood term in the model. Figures 8 and 9 display these results.

To demonstrate that super-linear cumulativity emerges from the MaxEnt model

¹¹Though in principle this same test could be carried out with NHG model, it is not clear what "success" would look like, given the apparent remoteness of the experimental results from the framework's parameter space.

		Condition				
Source	Phonotactics violated	A	В	С	D	E
Experiment 1	None	0.70	0.67	0.68	0.64	0.62
(for reference)	Backness	0.61	0.61	0.64	0.62	0.62
	Nasal	0.44	0.52	0.50	0.54	0.49
	Backness and nasal	0.35	0.40	0.38	0.44	0.40
Stochastic OT	None	1	1	1	1	1
	Backness	0.55	0.51	0.56	0.58	0.55
	Nasal	0.39	0.41	0.42	0.51	0.42
	Backness and nasal	0.30	0.39	0.32	0.38	0.32
Linear HG	None	1	1	1	1	1
	Backness	0.61	0.61	0.64	0.63	0.63
	Nasal	0.44	0.52	0.50	0.55	0.50
	Backness and nasal	0.35	0.39	0.39	0.42	0.38
Surface-based	None	0.60	0.61	0.61	0.60	0.58
	Backness	0.56	0.57	0.58	0.57	0.56
	Nasal	0.48	0.53	0.51	0.54	0.50
	Backness and nasal	0.44	0.49	0.48	0.52	0.48

Table 6: Mean group-level percent endorse for Experiment 1 (converted from predicted ratings by dividing by 100, obtained from the regression model in section 5), and best-fitting predicted values from Stochastic OT, Linear HG, and Surface-based grammars.

Model	Sum squared error	Mean squared error	Max squared error
MaxEnt (for reference)	0.0057	0.0003	0.0008
Stochastic OT	0.6438	0.0322	0.1444
Linear HG	0.5638	0.0212	0.1444
Surface-based	0.0765	0.0038	0.0105

Table 7: Sum, mean, and maximum squared error for the Stochastic OT, Linear HG, and Surfacebased models; the best-fitting MaxEnt model is repeated from table 5 for reference.

as discussed in section 3.1.1, we plot the predicted percent endorsement rate of the doubly-violating forms (purple) based on the product of the probability of their individual violations (*linear cumulativity*, as in Case 3 of figure 1) alongside the lower actual predicted endorsement rate under the model (blue).



Figure 8: MaxEnt predictions based on Experiment 1's training data.



Figure 9: Figure 8, zoomed to focus on the doubly-violating forms, demonstrating emergent super-linear cumulativity in the MaxEnt model.

8. General discussion

This paper sought to determine whether super-linear cumulativity observed in prior experimental work had its source in the synchronic grammar, or whether the data patterns ought to be explained via frequency matching a diachronicallyinfluenced lexicon. Examining the predictions of two constraint-based phonological frameworks, MaxEnt and NHG, we found that under certain conditions these frameworks exhibit a super-linear relationship between Harmony and probability. Further, these models predict that the linearity of *any* cumulative constraint interaction in the synchronic grammar should be sensitive to the strength of the constraints involved, operationalized here as relative exceptionfulness. We tested this prediction using an AGL experiment in which we systematically varied the number of exceptions to the two primary harmony phonotactics that learners were exposed to in training. We found that as the number of exceptions to each phonotactic in training increased, the well-formedness of the doublyviolating forms fell lower than predicted based on the linear sum of their penalties — that is, participants exhibited super-linear cumulativity of markedness violations, modulated by the strength of the constraints involved. Because the effect emerged in a controlled paradigm using an artificial language, we concluded that we must attribute the finding to the synchronic grammar, rather than simply to a side-effect of a diachronically-skewed lexicon. We also verified that this superlinear behavior was *emergent* from the interaction of the two constraints -aproperty of the grammar itself – rather than overtly learned from the training data.

Building on these experimental findings that implicated the synchronic grammar, we evaluated a range of probabilistic constraint-based phonological models on their ability to capture the data. We found that while a MaxEnt model was able to capture the participants' responses well, the NHG model fell short, a qualitatively similar pattern as the findings in Smith and Pater (2020). Further, as predicted, super-linear cumulativity emerged from the MaxEnt framework when exposed to the same exceptionful training data as the experimental participants.

8.1. The mathematical basis of super-linearity in MaxEnt

Because the MaxEnt framework is based on relatively simple equations relating Harmony and probability, we can explicitly outline the weighting conditions under which a MaxEnt grammar will depart from linear cumulativity. To start, we use variables to abstract away from specific constraint weights, and thus can characterize the probability of the singly-marked *poti*, violating AGR([back]), as being proportional to $e^{-w(AGR([back]))}$, and that of the other singly-marked type *ponu*, violating AGR([nas]), as being proportional to $e^{-w(AGR([nas]))}$. By the same logic, the probability of the doubly-marked type *poni*, violating both AGR([back]) and AGR([nas]), is proportional to $e^{-w(AGR([back]))} + e^{-w(AGR([nas]))}$, and the probability of the Null Parse is proportional to $e^{-w(MPARSE)}$.

It follows, therefore, that the probability of the backness-violating *poti*, when competing against the Null Parse, should be $\frac{e^{-w(AGR([back]))}}{Z}$, where Z is $e^{-w(AGR([back]))} + e^{-w(MPARSE)}$, and that the probability of nasal-violating *ponu* in its own competition against the Null Parse is $\frac{e^{-w(AGR([nas]))}}{Z}$, where Z is $e^{-w(AGR([nas]))} + e^{-w(MPARSE)}$. Continuing in this vein, the probability of the doubly-violating form *poni* in competition with the Null Parse is $\frac{e^{-w(AGR([nas]))}}{Z}$, where Z is $e^{-w(MPARSE)} + e^{-w(MPARSE)}$. Continuing in this vein, the probability of the doubly-violating form *poni* in competition with the Null Parse is $\frac{e^{-w(AGR([nas]))-w(AGR([back]))}}{Z}$, where Z is $e^{-w(MPARSE)} + e^{-w(AGR([back]))-w(AGR([nas]))}$. In contrast, we can can obtain the probability of violating both constraints by multiplying the probability of the forms with each of those individual violations. This is shown in equation 1.

$$\frac{e^{-w(\operatorname{Agr}([nas]))}}{e^{-w(\operatorname{Mgr}([nas]))} + e^{-w(\operatorname{Agr}([nas]))}} \times \frac{e^{-w(\operatorname{Agr}([back]))}}{e^{-w(\operatorname{Mgr}([back]))} + e^{-w(\operatorname{Agr}([back]))}}$$
(1)

This expression simplifies to the following:

$$\frac{e^{-w(\operatorname{Agr}([back]))-w(\operatorname{Agr}([nas]))}}{(e^{-w(\operatorname{MPARSE})} + e^{-w(\operatorname{Agr}([back]))})(e^{-w(\operatorname{MPARSE})} + e^{-w(\operatorname{Agr}([nas]))})}$$
(2)

This equation simplifies again, and allows us to characterize the joint probability of two markedness violations as the following:

$$\frac{e^{-w(\operatorname{Agr}([back]))-w(\operatorname{Agr}([nas]))}}{\left(e^{-w(\operatorname{MParse})-w(\operatorname{Agr}([back]))}+e^{-w(\operatorname{MParse})-w(\operatorname{Agr}([nas]))+}+e^{-w(\operatorname{Agr}([back]))}+e^{-2w(\operatorname{MParse})}\right)}$$
(3)

Comparing this quantity to the probability of the doubly-marked candidate in its own competition against the Null Parse, it becomes clear why certain weighting conditions in MaxEnt yields super-linear cumulativity: the denominators in equations 3 and 4 are not the same.

$$\frac{e^{-w(\operatorname{Agr}([back]))-w(\operatorname{Agr}([nas]))}}{e^{-w(\operatorname{MPARSE})} + e^{-w(\operatorname{Agr}([back]))-w(\operatorname{Agr}([nas]))}}$$
(4)

Because Harmony is computed via the simple addition of penalties before exponentiation in equation 4, the probability of the doubly-violating candidate *poni* is not guaranteed to equal the joint probability of the structures which make it marked (the backness harmony violation of *poti*, and the nasal harmony violation of *ponu*). We can examine the relationship between these two quantities by cancelling the $e^{-w(A_{GR}([back]))-w(A_{GR}([nas]))}$ term found in the numerator and denominator out of both equations 3 and 4, and compute the ratio of the remaining quantities to see exactly when MaxEnt will exhibit super-linear cumulativity of markedness violations.

$$\frac{e^{-w(\text{MPARSE})}}{e^{-w(\text{MPARSE})-w(\text{Agr}([back]))} + e^{-w(\text{MPARSE})-w(\text{Agr}([nas]))} + e^{-2w(\text{MPARSE})}}$$
(5)

If the ratio in equation 5 is greater than 1, MaxEnt will exhibit super-linear cumulativity of violations: the probability of the doubly-marked candidate will be less than the joint probability of the violating structures in the language. If it is less than 1, MaxEnt predicts sub-linear cumulativity: the probability of the doubly-marked candidate will be greater than the joint probability of the violating structures in the language. Finally, when the ratio is exactly 1 — that is, the denominators in equations 3 and 4 are equal — MaxEnt will exhibit linear cumulativity: the probability of the doubly-marked structure will equal the joint probability of its component structures. Since the weight of MPARSE is in both quantities in the ratio, the ratio is proportional to the weights of AGR([back]) and AGR([nas]): for any given weight of MPARSE, relatively higher weights for the two AGREE constraints will result in sub-linear cumulativity, and relatively lower weights will result in super-linear cumulativity, as stated in section 3.1.2.

8.2. Sub-linear cumulativity

To this point we have demonstrated experimentally that there is a dependency between the degree of strength of an individual phonotactic (operationalized here as exceptionality), and the type of cumulativity that it enters into. We have also demonstrated that a certain class of phonological theories – MaxEnt and to a lesser extent NHG – can capture the data well, and also exhibit a sigmoidal relationship between Harmony and probability (cf. figure 5). We have not demonstrated, however, that the non-linear relationship between Harmony and probability is *necessarily* sigmoidal in shape: all that is necessary to account for the experimental data is for the relationship between Harmony and probability to be some nonlinear function with a "shoulder" in it, resembling the upper half of the sigmoid curve drawn in figure 5 which gives rise to the super-linear cumulativity of weak markedness violations.

If we take the proposed theoretical mechanism at face value, however, the equations in section 8.1 suggest that will we see not only super-linear cumulativity of weak violations, but also *sub-linear cumulativity* of strong violations. Exactly such a finding comes from Pizzo (2015), who investigates the cumulative effect of violating exceptionless syllable-margin constraints in English, crossing an ill-formed onset with an ill-formed coda. She found that while each violation resulted in a significant decrement in assessed well-formedness, the penalty for a form containing both violations was less than that which is predicted by the summed penalty for each of the markedness violations. Though this finding certainly invites further study, it suggests that a sigmoidal relationship between Harmony and probability may be not only sufficient, but also necessary to describe the properties of the grammar. This finding corresponds to the predicted but heretofore undocumented Case 5 in figure 1.

8.3. Reconciling grammar and lexicon

In section 3.1.1 we considered the possibility that frequency-matching lexical statistics could explain the experimental findings of super-linearity by Albright (2012) and others. Although the experimental results presented in this paper suggest that this is not the case, the puzzle of super-linear underattestation in the lexicon remains unsolved. Albright's lexical study is not alone: a number of researchers have found evidence that some marked structures exhibit super-linear cumulativity in lexical counts, such that the probability of two marked structures calculated using the formula for linear cumulativity (Case 3, figure 1) is greater than the probability of actual lexical items containing both marked structures. Albright (2008) finds that Lakhota roots which contain multiple structures which are only moderately uncommon, such as consonant clusters and fricatives, cooccur in dramatically fewer roots than predicted by their joint probability. Also in this vein is a study by Yang et al. (2018), who carry out a comparison of English and Mandarin monosyllables and find that the attested sub-lexicons are more well-formed than would be expected by the joint probabilities of their parts. In a slightly different domain, Green and Davis (2014) find that multiple optional syllable structure simplifications in colloquial Bamana are dramatically less likely to co-occur than expected given the product of the probability of each independent simplification process. Kim (2019), building on Kumagai (2017), demonstrates the cumulative effect of nasals on blocking the inter-morpheme obstruent-voicing process *rendaku* in Japanese compounds which also displays super-linear behavior. Super-linear cumulativity has also been observed in the contribution of different phonological structures to the likelihood of belonging to a specific lexical class (Shih, 2017).

How can we reconcile these lexical findings with the synchronic mechanism demonstrated in this paper? Although it is tempting to extend the synchronic mechanism to account for the lexical data, doing so fails to take into account the diachronic origin of these lexical statistics (cf. Frisch (1996); Beguš (2016)). Although lexical statistics are often advanced as evidence of synchronic phonological knowledge, divergences between lexical statistics and productive grammatical knowledge are well-known (Hayes and White, 2013; Becker et al., 2011, among others). Thus simply observing that a generalization holds of a language's lexicon does not necessarily imply that it enjoys a cognitively real status in the synchronic grammar of its speakers. Further, while there is ample evidence that learners change their language over time (Martin, 2007), it is unlikely that such widespread under-representation can be the effect of a single generation of speakers. We suggest, then, that a full explanation of the lexical findings must take into account the bidirectional relationship between the synchronic phonological grammar and the diachronic trends which shape the lexicon. Such an undertaking is left for future research.

9. Conclusion

The primary contribution of this paper was to advance experimental evidence in favor of a synchronic phonological grammar in which the type of cumulativity constraints enter into (sub-linear, linear, or super-linear) is dependent on the strength of those constraints in the grammar. We have argued that MaxEnt, and to a lesser extent NHG, can capture the data successfully, and outlined the structure and predictions of the best-fitting MaxEnt model, including the conditions in which it does and does not predict super-linear cumulativity of violations. Finally, we adduced evidence of sub-linear cumulativity of strong markedness violations from the literature which further supports the specifically *sigmoidal* shape of the non-linear relationship between Harmony and probability that Max-Ent predicts.

The work presented here, however, is only a first step towards a fuller understanding of the empirical and typological landscape of non-linear cumulativity. The dependency between constraint strength and cumulative behavior proposed here makes strong predictions about both the wide scope of constraints that can enter into super-linearly cumulative relationships, and also specific claims about the weighting requirements that must be met for such effects to be observed. A great deal of further empirical research, using natural as well as artificial languages and lexicons, is therefore needed to test and refine these predictions going forward.

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